



Rainfall trends for the State of Paraná: present and future climate

ARTICLES doi:10.4136/ambi-agua.2258

Received: 21 Mar. 2018; Accepted: 18 Dec. 2018

Luciana Espindula de Quadros^{1*} ; **Eloy Lemos de Mello²** ;
Benedito Martins Gomes² ; **Fernanda Cristina Araujo²** 

¹Instituto Federal do Paraná (IFPR), Foz do Iguaçu, PR, Brasil

E-mail: luciana.quadros@ifpr.edu.br

²Universidade Estadual do Oeste do Paraná (UNIOESTE), Cascavel, PR, Brasil

E-mail: eloymello@gmail.com, benedito.gomes@unioeste.br, fer.crisaraujo@yahoo.com.br

*Corresponding author

ABSTRACT

This paper analyzes the variability and the precipitation trend of the State of Paraná, in Brazil. For that, monthly precipitation data belonging to 24 precipitation stations in a 30-year period (1980-2010) were analyzed and they were compared with projections of precipitation for the years 2016-2050. These data were simulated by Eta/Miroc5 for RCP 4.5 (Representative Concentration Pathways) from the Center for Weather Forecasting and Climate Studies CPTEC/INPE and the historical data of precipitation were taken from National Water Agency (ANA). The Mann-Kendall non-parametric test and the Sen's slope estimator were applied to detect trends and magnitudes, respectively. The Mann-Whitney test was used to compare the median of the historical series (1980-2010) with the simulated series (2016-2050) and the comparison of the means between the two series was performed by Test *t*. The results draw attention to the great variability and significant changes in the monthly average rainfall that may occur, if the climate change scenarios that were considered become a reality in the near future.

Keywords: climate changes, climate models, precipitation variability.

Tendência da precipitação no Estado do Paraná: clima presente e futuro

RESUMO

O objetivo deste artigo é analisar a variabilidade e a tendência de precipitação do Estado do Paraná, Brasil. Para isso, foram analisados os dados de precipitação mensal pertencentes a 24 estações de precipitação pluviométrica em um período de 30 anos (1980-2010) e foi comparado com as projeções futuras sobre a precipitação para os anos de 2016-2050. Estes dados foram simulados por Eta/Miroc5 para RCP 4.5 (Caminhos Representativos de Concentrações) proveniente do Centro de Previsão de Tempo e Estudos Climáticos CPTEC/INPE e os dados históricos de precipitação utilizados foram da Agência Nacional de Águas (ANA). O teste não-paramétrico de Mann-Kendall e o estimador de Sen foram aplicados para detectar tendências e sua magnitude, respectivamente. O teste de Mann-Whitney foi usado para comparar a mediana da série histórica (1980-2010) com a mediana da série simulada (2016-2050) e a comparação das médias entre as duas séries foram realizadas pelo Test *t*. Os resultados chamam atenção para a grande variabilidade e mudanças significativas na



precipitação média mensal que poderá ocorrer, se os cenários de mudanças climáticas que foram consideradas se tornarem uma realidade no futuro próximo.

Palavras-chave: modelos climáticos, mudanças climáticas, variabilidade de precipitação.

1. INTRODUCTION

Among the many effects that climate variability could be responsible for, a major concern is the imminence of a possible increase in the occurrence of extreme events around the globe, which could directly affect the human population and other living organisms. Several studies and investigations have been conducted in order to identify which elements are causing such impacts and what are their intensities and frequencies of occurrence (Zandonadi and Acquattro, 2016).

According to the Brazil Panel on Climate Change (PBMC, 2013), some regions of Brazil may experience changes in temperature and rainfall with global warming. Intensifications of severe events should occur, causing severe impacts in cities and areas vulnerable to climate change. The possible impacts of these changes will occur at different scales, according to the specific characteristics of each region of Brazil. It is necessary to know and map the vulnerabilities of Brazilian regions to identify, propose and implement adaptation measures.

Among all the variables, the precipitation is one of the most important meteorological variables which can impact the occurrence of drought or flood. Analysis of precipitation data yields relevant information which can be used to improve water management strategies, protect the environment, agricultural production planning or impact economic development of a certain region (Gocic and Trajkovic, 2013).

In recent years, scientists worldwide have compared and analyzed precipitation trends, but precipitation isn't the only focus; other meteorological parameters have also been studied for adaptation and mitigation strategies for potential future changes. The Global Climate Models (GCMs) and dynamical downscaling using Regional Climate Models (RCMs) are the main tools of the analyses aimed at assessing what climate we are likely to have in the near- and not-so-near future (Marengo *et al.*, 2012). According to Adam and Collischonn (2013), the GCMs take into account the behavior of climate compartments (atmosphere, oceans, vegetation, soils, etc.) and their interactions in a quantitative model (numerical), allowing the simulation of probable climate evolution projections for various scenarios of emissions of greenhouse gases. These simulations have resulted in four pathways being adopted by the IPCC, denominated "Representative Concentration Pathways" (RCPs). The four RCPs included one mitigation scenario leading to a very high baseline emission scenarios (RCP 8.5), two medium stabilization scenarios (RCP 4.5/RCP 6.0) and one very low forcing level (RCP 2.6) (Vuuren *et al.*, 2011).

The use of RCMs also contributed to information enhancement, because on a regional scale the information is more accurate and the computational advancement allowed the inclusion of more components and processes in the simulations and spatial resolution of the models, thus increasing their complexity (Marengo *et al.*, 2009). In South America, the Eta Model has been used to provide time forecasts, because this model is able to produce satisfactory results in regions that contain sharp orography, such as the Andean Mountains (Chou *et al.*, 2011).

Singh and Goyal (2016) analyzed the spatial and temporal variability of precipitation lapse rate and of precipitation extreme indices of the Teesta River catchment, which corresponds to north Sikkim eastern Himalayas. For this, they used series durations observed (1980–2005) and simulated precipitation data sets (2006–2100) simulated by CMIP5 ESM-2 M Model (Coupled Model Intercomparison Project Phase 5 Earth System Model 2), employing three different radiative forcing scenarios Representative Concentration Pathways (RCP).

In Korea, Kim *et al.* (2016) researched heat waves associated with climate change using a series of observational data from 1994 to 2012 to deduce the causal factors that affect the number of deaths from heat disorders, because the greater concern is that the steady increase in death rate is expected to be intercepted by more severe events in the future compared to the present period. To simulate future changes in heat wave incidences in Korea for the period from 2013 to 2060, they used data simulated using the Hadley Centre Global Environmental for two RCP scenarios: RCP 4.5 and RCP 8.5.

In the southeast Brazil, a study was done on possible changes in air temperature and the effect of this on coffee beverage quality. Based on climate projections using the Eta/HadCM3 for the period of 2011 to 2100, the simulation of future changes was possible. The effects of this process on coffee beverage quality were simulated and the results indicated that, in the case of an occurrence of A1B emission scenario, the coffee beverage quality could be affected in this region due to the fact that the flavor may become stronger and unpleasant caused by rising air temperatures (Giarolla *et al.*, 2012).

Therefore, this study sought to use simulations to understand changing trends in different contexts. This paper analyzes the variability and the precipitation trend of the observed precipitation data (1980-2010) and simulated data by Eta/Miroc5 (2016-2050) for RCP 4.5 of the State of Paraná, in Brazil.

2. MATERIAL AND METHODS

Monthly precipitation series from 1980 to 2010 from the National Water Agency (ANA) database from 24 precipitation stations in the State of Paraná were used. Figures 1 and 2 represent the spatial distribution of the stations. Data gaps were filled using values from the nearest neighboring observatories. To apply this method, two criteria were established: 1) the correlation (Pearson's r) between the monthly precipitation series from both observatories was required to be higher than 0.7 (Barbosa *et al.*, 2005; Pruski *et al.*, 2004); and 2) the nearest stations were considered, with a maximum distance of 100 km, taking into consideration that Baba *et al.* (2014) states that the nearest stations are more likely to have similar behaviors with respect to the treated variables (Table 1).

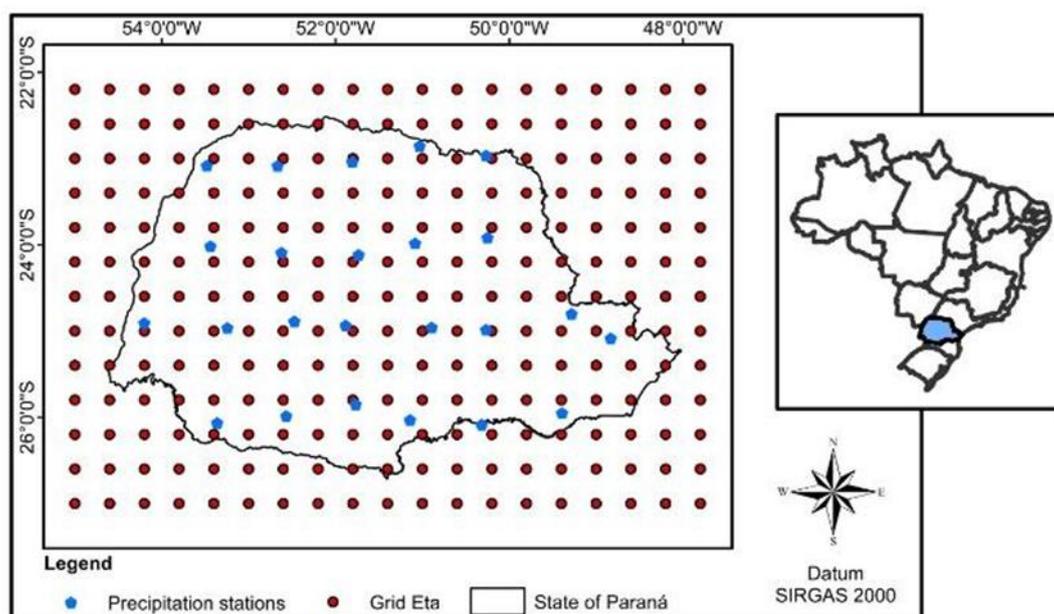


Figure 1. Spatial distribution of the 24 precipitation stations and grid of Eta/MIROC5 model with 824 pixels covering the State of Paraná.

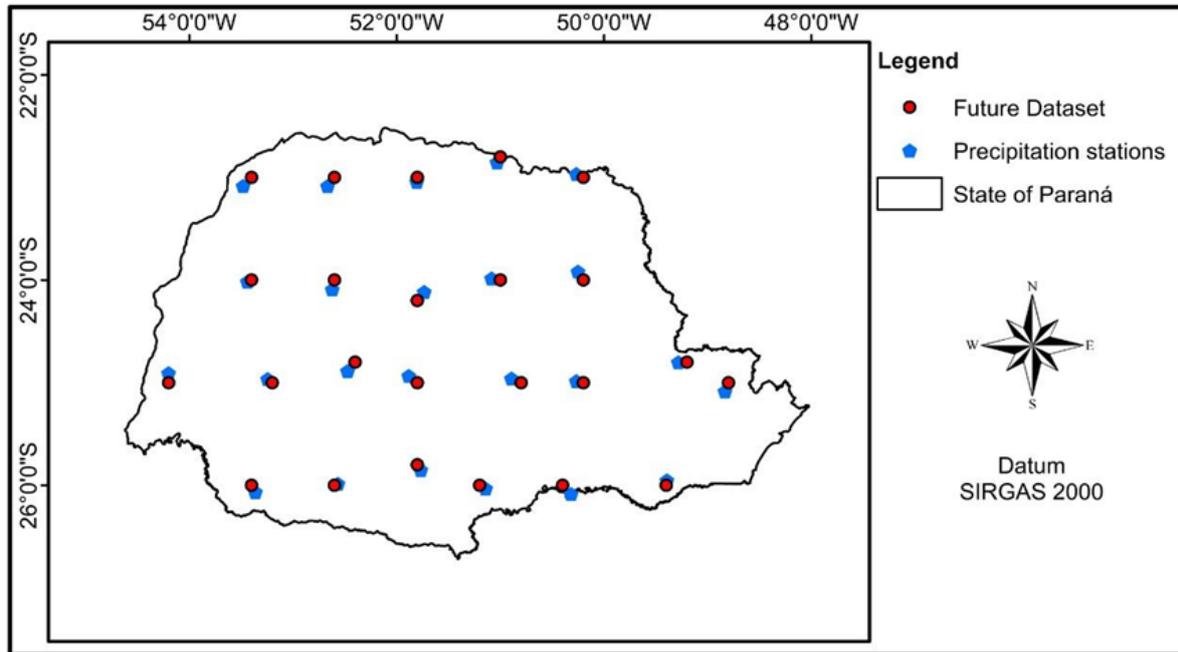


Figure 2. Spatial distribution of the 24 precipitation stations in the State of Paraná and the simulated data set simulated with Eta/MIROC.

In order to detect trends, the Mann-Kendall non-parametric statistical was used in the precipitation series. The series of 1980 to 2010 were put together with the series 2016 to 2050, considering a single series at each point studied. For this, the test statistic Z was used (Equations 1, 2, 3 and 4):

$$Z = \begin{cases} \frac{(S-1)}{\sqrt{\text{var}(S)}} & \text{if } S > 0 \\ 0 & \text{if } S = 0 \\ \frac{(S+1)}{\sqrt{\text{var}(S)}} & \text{if } S < 0 \end{cases} \quad (1)$$

where:

$$S = \sum_{i=1}^{n-1} \sum_{j=i+1}^n \text{sgn}(x_j - x_i) \quad (2)$$

$$\text{sgn}(x_j - x_i) = \begin{cases} +1, & \text{if } x_j > x_i \\ 0, & \text{if } x_j = x_i \\ -1, & \text{if } x_j < x_i \end{cases} \quad (3)$$

$$\text{var}(S) = \frac{n(n-1)(2n+5) - \sum_{i=1}^m t_i(t_i-1)(2t_i+5)}{18} \quad (4)$$

Where $\text{sgn}(x_j - x_i)$ is the sign function, n is the number of data points, m is the number of tied groups and t_i denotes the number of observations of the i -th tied group. A tied group is a set of sample data having the same value.

Positive values of Z show increasing trends while negative Z values indicate decreasing trends. For this study a 5% significance level was considered.

Table 1. The information of precipitation stations and simulated data set by RCP scenario 4.5 Miroc5.

Nº	Stations	Historical data			Simulated data set		Distance between them (km)
		Latitude	Longitude	Elevation	Latitude	Longitude	
1	2454018	-24.9061	-54.2014	243	-25.000	-54.2	10.41
2	2653016	-26.0644	-53.3622	557	-26.000	-53.4	8.08
3	2453056	-24.9628	-53.2439	697	-25.000	-53.2	6.05
4	2453008	-24.0147	-53.4397	427	-23.999	-53.4	4.39
5	2353010	-23.0817	-53.4811	349	-22.999	-53.4	12.37
6	2552044	-25.9833	-52.5667	700	-26.000	-52.6	3.81
7	2452019	-24.8858	-52.4739	741	-24.800	-52.4	12.06
8	2452029	-24.0906	-52.6214	582	-23.999	-52.6	10.37
9	2352052	-23.0831	-52.6667	450	-22.999	-52.6	11.55
10	2551019	-25.85	-51.7667	1245	-25.800	-51.8	6.46
11	2451010	-24.9333	-51.8833	900	-25.000	-51.8	11.19
12	2451049	-24.1144	-51.7342	618	-24.199	-51.8	11.5
13	2351065	-23.0403	-51.8056	485	-22.999	-51.8	4.6
14	2651005	-26.0306	-51.1419	840	-26.000	-51.199	6.64
15	2450049	-24.9575	-50.8917	743	-25.000	-50.799	10.47
16	2351041	-23.9831	-51.0831	1011	-23.999	-50.999	8.73
17	2251039	-22.8517	-51.0319	370	-22.799	-50.999	6.47
18	2650005	-26.0833	-50.3167	770	-26.000	-50.399	12.36
19	2450021	-24.9831	-50.2667	950	-25.000	-50.199	7.08
20	2350041	-23.9167	-50.25	600	-23.999	-50.199	10.48
21	2250033	-22.9653	-50.2661	423	-22.999	-50.199	7.82
22	2549003	-25.95	-49.3931	810	-26.000	-49.399	5.57
23	2449006	-24.8	-49.2833	270	-24.800	-49.199	8.52
24	2548040	-25.0833	-48.8333	670	-25.000	-48.799	9.86

The Mann-Kendall test allows us to detect statistically significant trends, but does not provide estimates of the steepness of the trend slope. For this reason, the application was complemented by Sen's slope estimator. It is defined as (Equation 5):

$$B = Md \left(\frac{x_i - x_j}{t_i - t_j} \right) \quad (5)$$

Where Md is the value of median and x_i and x_j are data values at times t_i and t_j ($i > j$), respectively.

The Mann-Kendall statistical test and Sen's slope estimator have often been used to detect trend in hydro-meteorological time series. (Gocic and Trajkovic, 2013; Santos and Frago, 2013; Yürekli, 2015).

To determine if there are any statistically significant differences between the medians, the Mann-Whitney test was applied; this the series of 1980 to 2010 was compared with the series 2016 to 2050. The Mann-Whitney test statistics are described as (Equation. 6):

$$Z = \frac{w - \frac{n \cdot m}{2}}{\sqrt{\frac{n \cdot m (n + m + 1)}{12}}} \quad (6)$$

Where n is the historic size series, m is the simulated size series, w is defined as the smaller value of w_1 and w_2 , where (Equations 7 and 8):

$$w_1 = r_1 - \frac{n(n+1)}{2} \quad (7)$$

$$w_2 = r_2 - \frac{m(m+1)}{2} \quad (8)$$

To calculate r_1 and r_2 , both groups (historic series and simulated series) were put in a single set. Afterwards, numeric ranks were assigned to all the observations. In case of equal numbers, an average of the ranking was made to find out their position. Then, the historical and simulated series were separated again, r_1 is the sum of the ranks of the historic series and r_2 is the sum of the ranks of the simulated series.

The null hypothesis is stated as the median of the historic series equals the median of the simulated series. If the p-value were less than or equal to the significance level of 0.05, the decision will be to reject the null hypothesis. It must be concluded that the difference between the population medians is statistically significant.

The 2-Sample Test t was used to compare the precipitation averages between two independent groups, the historical data and simulated data, and to determine if there is a significant difference between them. The null hypothesis is the difference between the population means ($\mu_1 - \mu_2$) equals zero. Rejecting the null hypothesis, it can be concluded that the difference between the population means that it is not equal to, or greater than, or less than the reference value chosen in case zero. The Test t (Equation 9):

$$t = \frac{\mu_A - \mu_B}{\sqrt{S^2 \left(\frac{1}{n_A} + \frac{1}{n_B} \right)}} \quad \text{where } S^2 = \frac{(n_A - 1)S_A^2 + (n_B - 1)S_B^2}{n_A + n_B - 2} \quad (9)$$

In which: S_A^2 and S_B^2 are the variances of precipitation values of historical and simulated data; μ_A and μ_B sample means of precipitation values of historical and simulated data, n_A and n_B is the length of the series, respectively. If the variances are equal, S^2 will correspond to the variance of both samples.

3. RESULTS AND DISCUSSION

For the period 1980 to 2010 and 2016 to 2050, monthly trends of precipitation were obtained by the Mann-Kendall test, the Sen's slope estimator, the Mann-Whitney Test and the Test t and the results are given in Table 2.

According the results of the Mann Kendall test, there is a significant trend in monthly precipitation series and this was detected on seventeen points of the twenty-four points at the 5% significance level, while other points had no significant trends.

The monthly precipitation at fifteen points over the State of Paraná tends to increase until 2050 and two points tend to decrease. The points with increased trend are 6, 7, 9, 10, 12-15, 17-21, 23 and 24. The points with decreased trend are point 2 and point 11 (Table 2).

From 1980 to 2050, the monthly rainfall has been declining at the rate of 0.0605 mm month⁻¹ at point 2 and 0.0369 mm month⁻¹ at point 11. The points that have an increasing trend, the rates of raising are 0.0967, 0.1307, 0.0479, 0.1273, 0.1642, 0.0515, 0.1186, 0.1641, 0.0018, 0.1612, 0.1372, 0.0443, 0.0319, 0.0428, 0.0192, 0.1896 and 0.0794 mm month⁻¹, respectively in points 6, 7, 9, 10, 12, 13, 14, 15, 17, 18, 19, 20, 21, 23 and 24.

Still in Table 2, note the results of the Mann-Whitney test (nonparametric) and t tests (parametric). The test results relate to the use of the statistics at each point comparing the period of 1980 to 2010 and 2016 to 2050. It is observed that, even as with the Mann-Kendall test, seventeen points of the twenty-four points show a trend with a significance level of 5%, meaning that there is a change in the trend of monthly precipitation in the State of Paraná. Only at point 20 does the Mann-Whitney test differ from the other tests, but at the other points, the Mann-Whitney test and Test t confirm the result of the Mann-Kendall test.

Table 2. Results of statistical tests trend of rainfall in the State of Paraná.

	Mann-Kendall		Sen's slope		Mann-Whitney		Test <i>t</i>		
	<i>p-value</i> increasing	<i>p-value</i> decreasing	<i>mm</i> <i>month⁻¹</i>		Median	<i>p-value</i>	Mean	Increment	<i>p-value</i>
1	0.501	0.498	0.0000	Hist1	130.20		142.8		
				Fut 1	121.90	0.992	143.4	+0.42%	0.930
2	0.999	0.000*	-0.0605	Hist2	160.80		177		
				Fut 2	120.77	0.000*	142.2	-19.66%	0.000*
3	0.083	0.917	0.0218	Hist3	153.10		168		
				Fut 3	154.92	0.204	178	+5.95%	0.195
4	0.243	0.757	0.0089	Hist4	122.65		136.6		
				Fut 4	121.91	0.584	140.1	+2.56%	0.590
5	0.153	0.846	0.0137	Hist5	115.00		136.8		
				Fut 5	121.00	0.256	143.9	+5.19%	0.295
6	0.000*	1.000	0.0967	Hist6	153.15		169		
				Fut 6	207.08	0.000*	227	+34.32%	0.000*
7	0.000*	1.000	0.1307	Hist7	138.70		157		
				Fut 7	204.38	0.000*	230	+46.49%	0.000*
8	0.399	0.600	0.0031	Hist8	122.30		138.9		
				Fut 8	121.37	0.963	140.3	+1.01%	0.821
9	0.000*	0.999	0.0479	Hist9	100.55		116.1		
				Fut 9	121.77	0.000*	141.3	+21.7%	0.000*
10	0.000*	1.000	0.1273	Hist10	167.35		180		
				Fut 10	229.32	0.000*	251	+39.44%	0.000*
11	0.995	0.004*	-0.0369	Hist11	144.70		164.7		
				Fut 11	122.00	0.001*	144.1	-12.51%	0.003*
12	0.000*	1.000	0.1642	Hist12	119.75		131.5		
				Fut 12	197.46	0.000*	226	+71.86%	0.000*
13	0.000*	0.999	0.0515	Hist13	103.10		120.4		
				Fut 13	128.94	0.000*	148.6	+23.42%	0.000*
14	0.000*	1.000	0.1186	Hist14	134.15		148.8		
				Fut 14	201.88	0.000*	218	+46.5%	0.000*
15	0.000*	1.000	0.1641	Hist15	122.10		139.5		
				Fut 15	212.52	0.000*	235	+68.46%	0.000*
16	0.451	0.549	0.0018	Hist16	133.20		151		
				Fut 16	121.85	0.508	147	-2.65%	0.551
17	0.000*	1.000	0.1612	Hist17	93.55		114.5		
				Fut 17	183.20	0.000*	206	+79.91%	0.000*
18	0.000*	1.000	0.1372	Hist18	112.65		128.4		
				Fut 18	193.35	0.000*	202	+57.32%	0.000*
19	0.000*	0.999	0.0443	Hist19	110.65		123.5		
				Fut 19	121.61	0.011*	142.6	+15.46%	0.003*
20	0.009*	0.991	0.0319	Hist20	112.70		129.1		
				Fut 20	119.89	0.063	142.8	+10.61%	0.043*
21	0.000*	0.999	0.0428	Hist21	100.80		116.9		
				Fut 21	121.83	0.001*	139.3	+19.16%	0.000*
22	0.064	0.935	0.0192	Hist22	120.50		129.4		
				Fut 22	121.38	0.514	137.9	+6.57%	0.164
23	0.000*	1.000	0.1896	Hist23	101.15		113.8		
				Fut 23	197.23	0.000*	216	+89.81%	0.000*
24	0.000*	1.000	0.0794	Hist24	115.70		129.0		
				Fut 24	152.82	0.000*	164.3	+27.36%	0.000*

* Statistically significant trends at the 5% significance level.

A study by Cera and Ferraz (2015) analyzed precipitation trends in Rio Grande do Sul for three sets of data, with a comparison of the observed precipitation data (1982-2006) and simulated data by RegCM3 Regional Model (2070-2086) also showed an increasing trend for that State. They applied the Run test and Pettit test to certify the homogeneity of data for the series 1982 to 2006; but the tests gave different results and the Mann-Kendall test was used and this confirmed the Pettit test, indicating an increased trend of precipitation at a significant level of 99% in extreme north and central areas of the State of Rio Grande do Sul. The same study also analyzed the trend for future climate and the increase in precipitation trend at a 90% level of significance in the western, southern and central state was observed.

Other research similar to the study above was that of Doyle *et al.* (2012) in the La Plata Basin (LPB), where the northern part of this basin includes the State of Paraná. They employed the PRECIS regional climate modeling system for the annual rainfall and used precipitation data in the recent past (1961–1990) and in a future (2071–2100) climate at 5 and 10% confidence levels and all stations showed positive trends with the values exceeding 8 mm yr⁻¹ over southern Brazil. For that same location, Marengo *et al.* (2012) made use of the Eta/HadCM3 Model configured with a 40-km grid size and was run over 1961 to 1990 to represent baseline climate, and 2011 to 2100 to simulate possible future changes. They detected a rainfall increase in the South East of South America (by about 30–50%).

According to PBMC (2013), in southern Brazil and northern Argentina, trends for increased rainfall and river flows have been observed since the mid-twentieth century, with Paraná and La Plata Rivers showing a downward trend from 1901 to 1970 and a systematic increase in flows from the early 1970s to the present. It is also claimed that Southern region disasters such as landslides, death by drowning and building collapses may be more frequent. Agriculture and livestock in the region are vulnerable to climate change. There are numerous records of the El Niño Southern Oscillation (ENSO) phenomenon, which accentuates the conditioning/adversity features of the climate on agricultural production, determining production records or widespread losses. It should be added to the weather condition already affected by the ENSO, records of air temperature increase in various municipalities of Rio Grande do Sul, Santa Catarina and Paraná, which inevitably influence not only the agriculture and food security, but environmental conditions such as the hydrologic cycle and population health. Increases in rainfall and river flow, despite the uncertainties, are likely to intensify, as projected by the IPCC scenarios (IPCC, 2014). Temperatures will follow the pattern of increase in average values, with reduced episodes of frosts and cold days.

The location of points and the direction of trends for monthly precipitation data series during the period 1980 to 2050 are shown in Figure 3. Significant increasing trends were found in almost all the Mesoregions for the State of Paraná, except in West and West Central, where there wasn't any trend. In the South West and South Central, two points were detected for decreasing trend, but increasing trends also were found. This result is similar to that obtained by Zandonadi and Acquavotta (2016) for Paraná State: both results exhibited an increasing trend in the region of Curitiba (located in Metropolitan Curitiba), Castro (located in East Central), Irati (located in South East), Maringá and Londrina (both located in North Central), but differ in Campo Mourão region (located in West Central), where this study didn't show a trend and the other showed decreased trend.

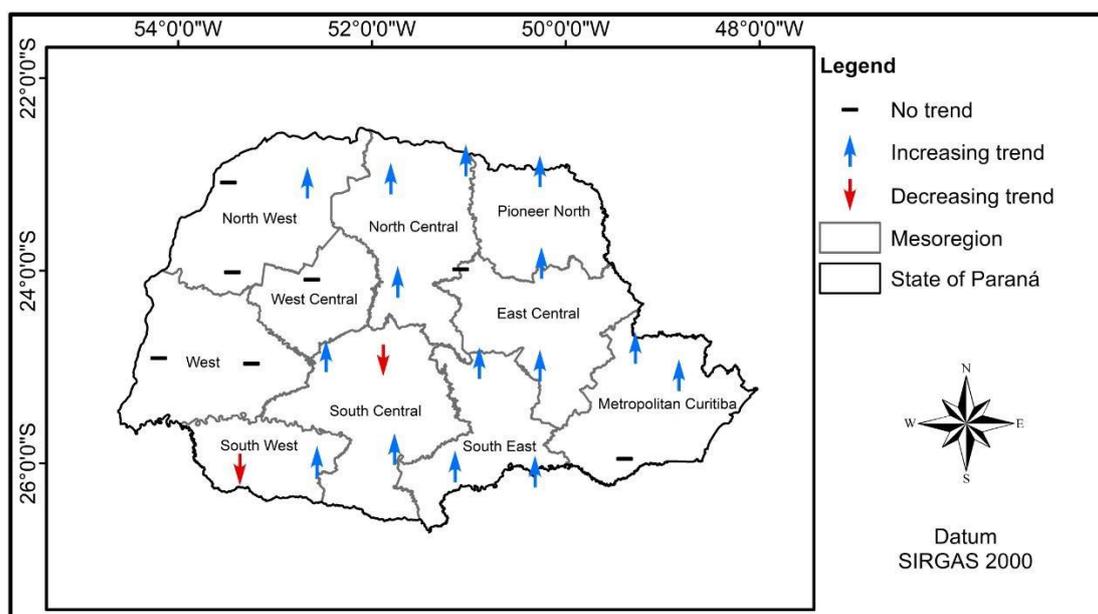


Figure 3. Spatial distribution of precipitation stations with increasing, decreasing and no trends during the period 1980 to 2050.

The Figure 4 shows the monthly precipitation of seventeen points where there is trend increasing and decreasing the values. A large variability in monthly precipitation accumulation in graphics is observed, for instance, in Figure 4j, where the maximum value of monthly accumulation reaches approximately 1000 mm, whereas the graphic of Figure 4o) reaches 500 mm/month⁻¹. The red line confirms the trend displayed in Table 2 and adrees with Grimm *et al.* (2000), that say that in South Brazil there are seasonal characteristics of rainfall that can change when there is the El Niño/La Niña, causing extreme precipitation events as excess or deficit.

Observe the high variability between each month in Figure 4. In Figure 4b), the average of the historical series is 169 mm month⁻¹ and the mean in the simulated series is 227 mm month⁻¹. This represents an increase of 34.32%. In the time series of graphic Figure 4k), the means are 114.5 and 206 mm month⁻¹, an increase of 79.91%. The same happens in the graphic of Figure 4l), a rain increase of 57.32%. According to Table 2, the maximum increment (89.81%) is in Figure 4p), and the maximum reduction (19.66%) in Figure 4a). Thus, if materializing expectations of these increases, there may be positive and negative impacts. An example of a positive impact is the replenishment of groundwater and an example of a negative impact is the increase of floods.

4. CONCLUSIONS

Considering the historical series (1980-2010) with the simulated series (2016-2050) that was simulated by the Eta/MIROC5 at twenty-four points, monthly precipitation showed increasing trends at fifteen locations in the State of Paraná and two sites showed a decrease in the monthly rainfall trend.

In some places, there is a big difference between the average monthly precipitation of 1980 to 2010 compared with the average expected precipitation of 2016 to 2050.

The results also exhibited great variability in monthly rainfall for both the historical series and the simulated series.

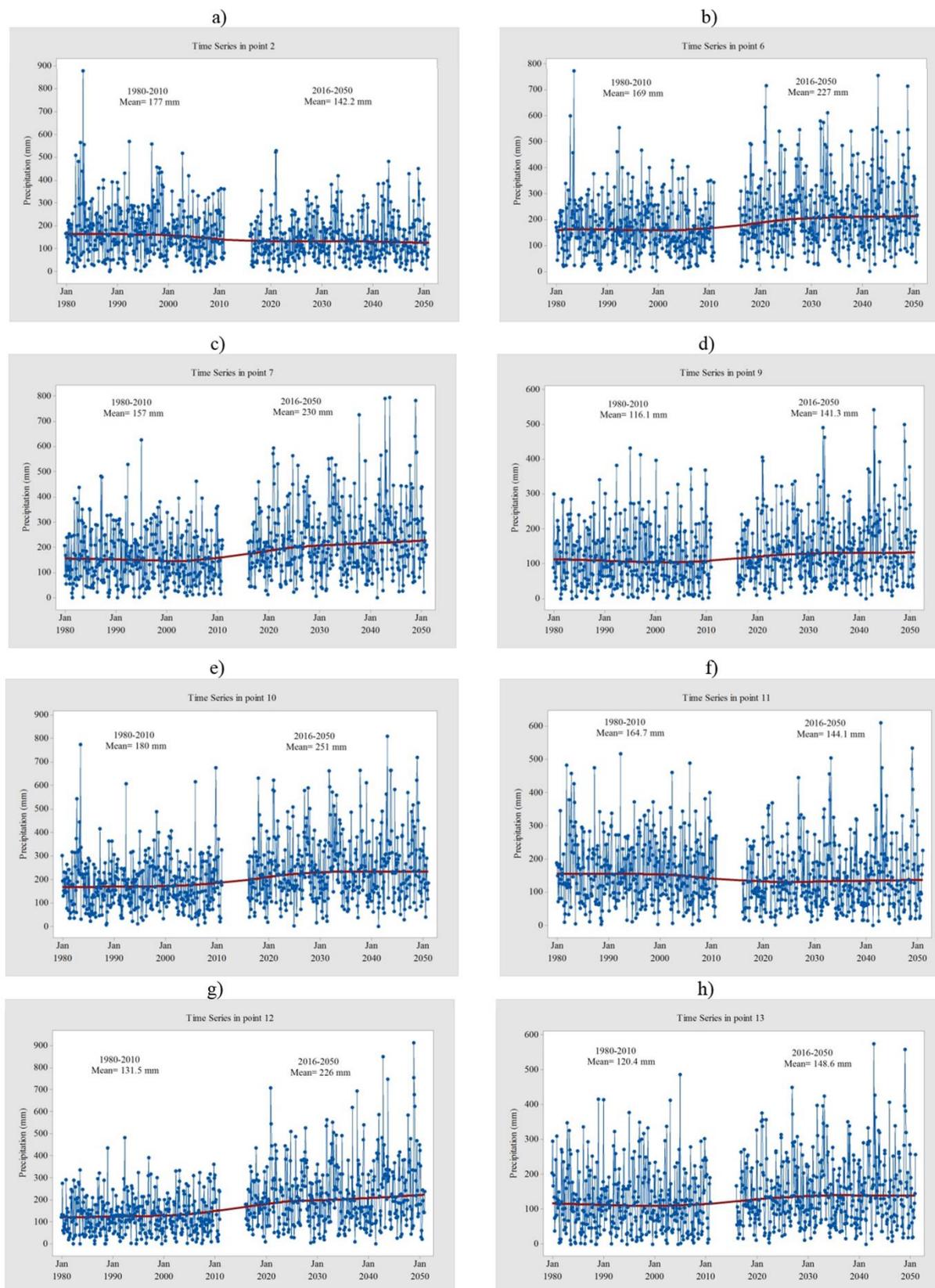


Figure 4. Temporal variation at points where show changes in precipitation trends during the period 1980-2050, along with the mean of historical and simulated precipitation. Continue.

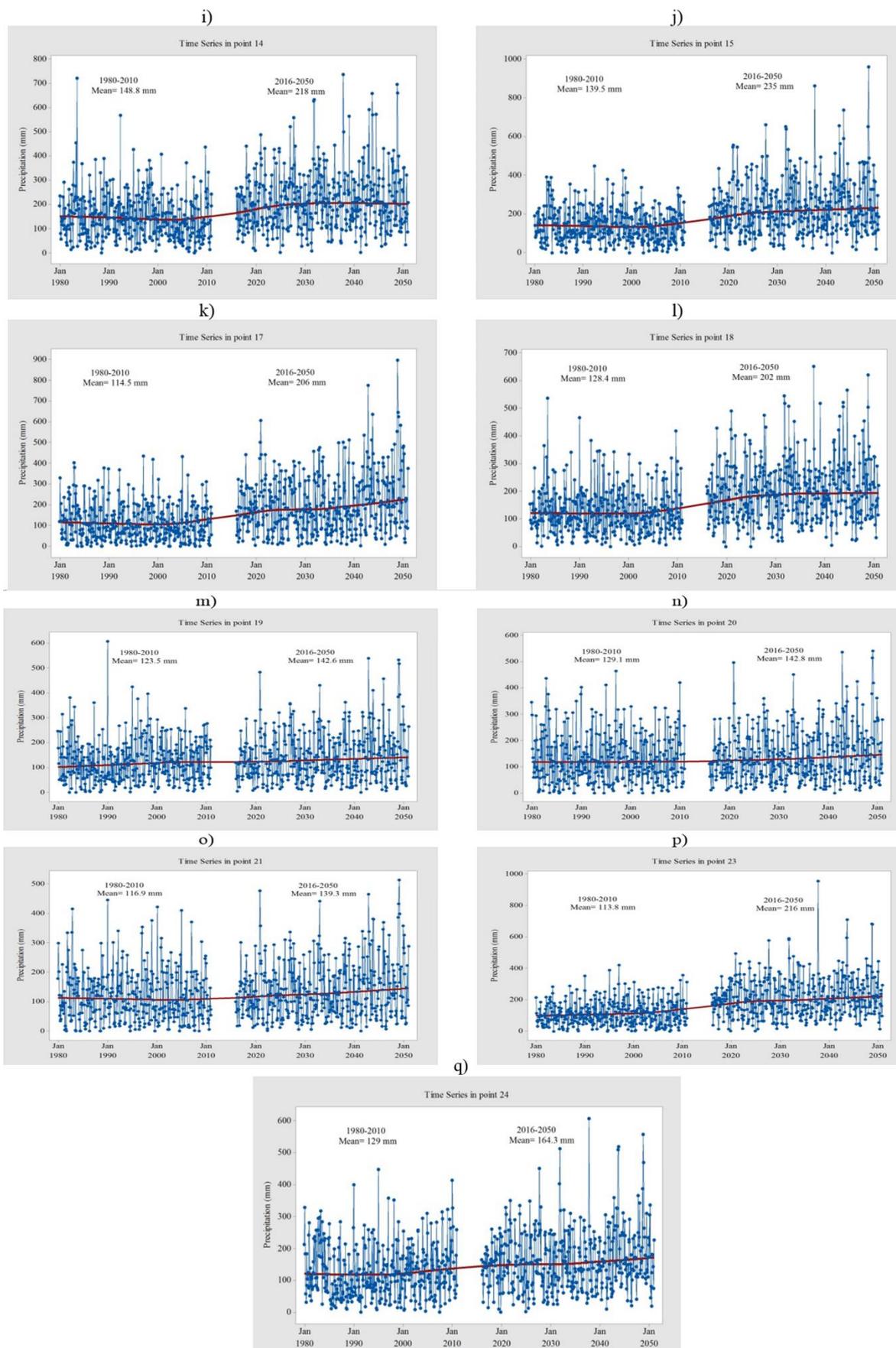


Figure 4. Continued.

5. REFERENCES

- ADAM, K. N.; COLLISCHONN, W. Análise dos impactos de mudanças climáticas nos regimes de precipitação e vazão na bacia hidrográfica do rio Ibicuí. **Revista Brasileira de Recursos Hídricos**, v. 18, n. 3, p. 69-79, 2013. <http://dx.doi.org/10.21168/rbrh.v18n3.p69-79>
- BABA, R. K.; VAZ, M. S. M. G.; COSTA, J. D. Correção de dados agrometeorológicos utilizando métodos estatísticos. **Revista Brasileira de Meteorologia**, v. 29, n.4, p. 515–526, 2014. <http://dx.doi.org/10.1590/0102-778620130611>
- BARBOSA, S. E. S.; BARBOSA JÚNIOR, A. R.; SILVA, G. Q.; CAMPOS, E. N. B.; RODRIGUES, V.C. Geração de modelos de regionalização de vazões máximas, médias de longo período e mínimas de sete dias para a Bacia do Rio do Carmo, Minas Gerais. **Revista Engenharia Sanitária e Ambiental**, v. 10, n. 1, p. 64-71, 2005. <http://dx.doi.org/10.1590/S1413-41522005000100008>
- CERA, J. C.; FERRAZ, S. E. T. Variações climáticas na precipitação no sul do Brasil no clima presente e futuro. **Revista Brasileira de Meteorologia**, v. 30, n.1, p. 81–88, 2015. <http://dx.doi.org/10.1590/0102-778620130588>
- CHOU, S. C.; MARENGO, J. A.; LYRA, A. A.; SUEIRO, G.; PESQUERO, J. F.; ALVES, L. M.; KAY, G.; BETTS, R.; CHAGAS, D. J.; GOMES, J. L.; BUSTAMANTE, J. F.; TAVARES, P. Downscaling of South America present climate driven by 4-member HadCM3 runs. **Journal Climate Dynamics**, v. 38, p. 635-653, 2011. <http://dx.doi.org/10.1007/s00382-011-1002-8>
- DOYLE, M. E.; SAURRAL, R. I.; BARROS, V. R. Trends in the distributions of aggregated monthly precipitation over the La Plata Basin. **International Journal of Climatology**, v. 32, n. 14, p. 2149–2162, 2012. <http://dx.doi.org/10.1002/joc.2429>
- GIAROLLA, A.; RESENDE, N.; CHOU, S. C.; TAVARES, P. S.; RODRIGUES, D. C. Coffee beverage quality assessment based on Eta/CPTEC-HadCM3 model (A1B-IPCC/SRES scenario), Southeastern Brazil. **Geophysical Research Abstracts**, v. 14, n. 1, p. 6224, 2012.
- GOCIC, M.; TRAJKOVIC, S. Analysis of changes in meteorological variables using Mann-Kendall and Sen's slope estimator statistical tests in Serbia. **Global and Planetary Change**, v.100, p. 172–182, 2013. <http://dx.doi.org/10.1016/j.gloplacha.2012.10.014>
- GRIMM, A. M.; BARROS, V. R.; DOYLE, M. E. Climate Variability in Southern South America Associated with El Niño and La Niña Events. **Journal of Climate**, v. 13, p. 35–58, 2000. [https://doi.org/10.1175/1520-0442\(2000\)013%3C0035:CVISSA%3E2.0.CO;2](https://doi.org/10.1175/1520-0442(2000)013%3C0035:CVISSA%3E2.0.CO;2)
- IPCC. **Climate change 2014: mitigation of climate change**. Contribution of working group III. Cambridge: Cambridge University Press, 2014.
- KIM, D. W.; DEO, R. C.; CHUNG, J. H.; LEE, J. S. Projection of heat wave mortality related to climate change in Korea. **Natural Hazards**, v. 80, p. 623-637, 2016. <http://dx.doi.org/10.1007/s11069-015-1987-0>

- MARENGO, J. A.; HOU, S. C.; KAY, G.; ALVES, L. M.; PESQUERO, J. F.; SOARES, W. R.; SANTOS, D. C.; LYRA, A. A.; SUEIRO, G.; BETTS, R.; CHAGAS, D. J.; GOMES, J. L.; BUSTAMANTE, J.; TAVARES, P. Development of regional future climate change scenarios in South America using the Eta CPTEC/HadCM3 climate change projections: climatology and regional analyses for the Amazon, São Francisco and the Paraná River basins. **Climate Dynamics**, v. 38, p. 1829-1848, 2012. <http://dx.doi.org/10.1007/s00382-011-1155-5>
- MARENGO, J. A.; JONES, R.; ALVES, L. M.; VALVERDE, M. C. Future Change of Temperature and Precipitation Extremes in South America as Derived from the PRECIS Regional Climate Modeling System. **International Journal of Climatology**, v. 29, n. 15, p. 2241-2255, 2009. <http://dx.doi.org/10.1002/joc.1863>
- PBMC. **Executive Summary**: Impacts, vulnerability and adaptation to climate change. Contribution from GT2 to the Primeiro Relatório de Avaliação Nacional sobre de Mudanças Climáticas (RAN1) of the painel brasileiro de mudança climática (PBMC). Rio de Janeiro: UFRJ, 2013. 28p.
- PRUSKI, F. F.; PEREIRA, S. B.; NOVAES, L. F.; Silva, D. D.; Ramos, M. M. Precipitação média anual e vazão específica média de longa duração, na Bacia do São Francisco. **Revista Brasileira de Engenharia Agrícola e Ambiental**, v. 8, n. 2/3, p.2 47-253, 2004. <http://dx.doi.org/10.1590/S1415-43662004000200013>
- SANTOS, M.; FRAGOSO, M. Precipitation variability in Northern Portugal: data homogeneity assessment and trends in extreme precipitation indices. **Atmospheric Research**, v. 131, p. 34-45, 2013. <http://dx.doi.org/10.1016/j.atmosres.2013.04.008>
- SINGH, V.; GOYAL, M. K. Analysis and trends of precipitation lapse rate and extreme indices over north Sikkim eastern Himalayas under CMIP5ESM-2M RCPs experiments. **Atmospheric Research**, v. 167, p. 34-60, 2016. <http://dx.doi.org/10.1016/j.atmosres.2015.07.005>
- VUUREN, D P. V.; EDMONDS, J.; KAINUMA, M.; RIAHI, K.; THOMSON, A.; HIBBARD, K.; HURTT, G. C.; KRAM, T.; KREY, V.; LAMARQUE, J. F.; MASUI, T.; MEINSHAUSEN, M.; NAKICENOVIC, N.; SMITH, S. J.; ROSE, S. K. The representative concentration pathways: an overview. **Climatic Change**, v. 109, p. 5-31, 2011. <http://dx.doi.org/10.1007/s10584-011-0148-z>
- YÜREKLI, K. Impact of climate variability on precipitation in the Upper Euphrates – Tigris Rivers Basin of Southeast Turkey. **Atmospheric Research**, v. 154, p. 25-38, 2015. <http://dx.doi.org/10.1016/j.atmosres.2014.11.002>
- ZANDONADI, L.; ACQUAOTTA, F.; FRATIANNI, S.; ZAVATTINI, J. A. Changes in precipitation extremes in Brazil (Paraná River Basin). **Theoretical and Applied Climatology**, v. 123, p. 741–756, 2016. <http://dx.doi.org/10.1007/s00704-015-1391-4>